

The limit distribution of ratios of jumps and sums of jumps of subordinators

Péter Kevei

MTA-SZTE Analysis and Stochastics Research Group
Bolyai Institute, Aradi vértanúk tere 1, 6720 Szeged, Hungary
e-mail: kevei@math.u-szeged.hu

David M. Mason

Department of Applied Economics and Statistics
University of Delaware
213 Townsend Hall, Newark, DE 19716, USA
e-mail: davidm@udel.edu

December 2, 2014

Abstract

Let V_t be a driftless subordinator, and let denote $m_t^{(1)} \geq m_t^{(2)} \geq \dots$ its jump sequence on interval $[0, t]$. Put $V_t^{(k)} = V_t - m_t^{(1)} - \dots - m_t^{(k)}$ for the k -trimmed subordinator. In this note we characterize under what conditions the limiting distribution of the ratios $V_t^{(k)}/m_t^{(k+1)}$ and $m_t^{(k+1)}/m_t^{(k)}$ exist, as $t \downarrow 0$ or $t \rightarrow \infty$.

Keywords: Subordinator, Jump sequence, Lévy process, Regular variation, Tauberian theorem.

MSC2010: 60G51, 60F05.

1 Introduction and results

Let V_t , $t \geq 0$, be a subordinator with Lévy measure Λ and drift 0. Its Laplace transform is given by

$$\mathbf{E}e^{-\lambda V_t} = \exp \left\{ -t \int_0^\infty (1 - e^{-\lambda v}) \Lambda(dv) \right\},$$

where the Lévy measure Λ satisfies

$$\int_0^\infty \min\{1, x\} \Lambda(dx) < \infty. \quad (1)$$

Put $\bar{\Lambda}(x) = \Lambda((x, \infty))$. Then $\bar{\Lambda}(x)$ is nonincreasing and right continuous on $(0, \infty)$. When $t \downarrow 0$ we also assume that $\bar{\Lambda}(0+) = \infty$, which is necessary and sufficient to assure that there is an infinite number of jumps up to time t , for any $t > 0$.

Denote $m_t^{(1)} \geq m_t^{(2)} \geq \dots$ the ordered jumps of V_s up to time t , and for $k \geq 0$ consider the trimmed subordinator

$$V_t^{(k)} = V_t - \sum_{j=1}^k m_t^{(j)}.$$

We investigate the asymptotic distribution of jump sizes as $t \downarrow 0$ and $t \rightarrow \infty$. Specifically, we shall determine a necessary and sufficient condition in terms of the Lévy measure Λ for the convergence in distribution of the ratios $V_t^{(k)}/m_t^{(k+1)}$ and $m_t^{(k+1)}/m_t^{(k)}$. Observe in this notation that $V_t^{(0)} = V_t$ is the subordinator and $m_t^{(1)}$ is the largest jump.

An extended random variable W can take the value ∞ with positive probability, in which case W has a defective distribution function F , meaning that $F(\infty) < 1$. We shall call an extended random variable proper, if it is finite a.s. In this case its F is a probability distribution, i.e. $F(\infty) = 1$. Here we are using the language of the definition given on p. 127 of Feller [8].

Theorem 1. *For any choice of $k \geq 0$ the ratio $V_t^{(k)}/m_t^{(k+1)}$ converges in distribution to an extended random variable W_k as $t \downarrow 0$ ($t \rightarrow \infty$) if and only if one of the following holds:*

- (i) $\bar{\Lambda}$ is regularly varying at 0 (∞) with parameter $-\alpha$, $\alpha \in (0, 1)$, in which case W_k is a proper random variable with Laplace transform

$$g_k(\lambda) = \frac{e^{-\lambda}}{\left[1 + \alpha \int_0^1 (1 - e^{-\lambda y}) y^{-\alpha-1} dy\right]^{k+1}}; \quad (2)$$

- (ii) $\bar{\Lambda}$ is slowly varying at 0 (∞), in which case $W_k = 1$ a.s.;

- (iii) the condition

$$\frac{x\bar{\Lambda}(x)}{\int_0^x u\Lambda(du)} \longrightarrow 0 \quad \text{as } x \downarrow 0 \text{ (} x \rightarrow \infty \text{)} \quad (3)$$

holds, in which case $V_t^{(k)}/m_t^{(k+1)} \xrightarrow{\mathbf{P}} \infty$, that is $W_k = \infty$ a.s.

Note that Theorem 1 says that the situation $0 < \mathbf{P}\{W_k = \infty\} < 1$ cannot happen.

The corresponding problem for nonnegative i.i.d. random variables was investigated by Darling [6] and Breiman [4], in the $k = 0$ case. In this case Darling proved the sufficiency parts corresponding to (i) and (ii) (Theorem 5.1 and Theorem 3.2 in [6]), in particular the limit W_0 has the same distribution as given by Darling in his Theorem 5.1, while Breiman proved the necessity parts corresponding to (i), (ii) and (iii) (Theorem 3 (p. 357), Theorem 2 and Theorem 4 in [4]). A special case of Theorem 1 in Teugels [12] gives the sufficiency analog of (i) in the case of i.i.d. nonnegative sums for any $k \geq 0$.

The necessary and sufficient condition in the cases (ii) and (iii), stated in the more general setup of Lévy processes without a normal component, is given by Buchmann, Fan and Maller [5], see their Theorem 3.1 and 5.1.

Next we shall investigate the asymptotic distribution of the ratio of two consecutive ordered jumps $m_t^{(k+1)}/m_t^{(k)}$, $k \geq 1$. We shall obtain the analog for subordinators of a special case of a result that Bingham and Teugels [3] established for i.i.d. nonnegative random variables. This will follow from a general result on the asymptotic distribution of ratios of the form defined for $k \geq 1$ by

$$r_k(t) = \frac{\psi(S_{k+1}/t)}{\psi(S_k/t)}, \quad t > 0,$$

where for each $k \geq 1$, $S_k = \omega_1 + \dots + \omega_k$, with $\omega_1, \omega_2, \dots$ being i.i.d. mean 1 exponential random variables and ψ is the nonincreasing and right continuous function defined for $s > 0$ by

$$\psi(s) = \sup\{y : \bar{\Pi}(y) > s\},$$

with Π being a positive measure on $(0, \infty)$ such that $\bar{\Pi}(x) = \Pi((x, \infty)) \rightarrow 0$, as $x \rightarrow \infty$. Note that we do not require Π to be a Lévy measure. Also whenever we consider the asymptotic distribution of $r_k(t)$ as $t \downarrow 0$ we shall assume that $\bar{\Pi}(0+) = \infty$.

We call a function f *rapidly varying at 0* with index $-\infty$, $f \in \text{RV}_0(-\infty)$, if

$$\lim_{x \downarrow 0} \frac{f(\lambda x)}{f(x)} = \begin{cases} 0, & \text{for } \lambda > 1, \\ 1, & \text{for } \lambda = 1, \\ \infty, & \text{for } \lambda < 1. \end{cases}$$

Correspondingly, a function f is *rapidly varying at ∞* with index $-\infty$, $f \in \text{RV}_\infty(-\infty)$, if the same holds with $x \rightarrow \infty$.

Theorem 2. *For any choice of $k \geq 1$ the ratio $r_k(t)$ converges in distribution as $t \downarrow 0$ ($t \rightarrow \infty$) to a random variable Y_k if and only if one of the following holds:*

- (i) $\bar{\Pi}$ is regularly varying at 0 (∞) with parameter $-\alpha \in (-\infty, 0)$, in which case Y_k has the $\text{Beta}(k\alpha, 1)$ distribution, i.e.

$$G_k(x) = \mathbf{P}\{Y_k \leq x\} = x^{k\alpha}, \quad x \in [0, 1]; \quad (4)$$

- (ii) $\bar{\Pi}$ is slowly varying at 0 (∞), in which case $Y_k = 0$ a.s.

- (iii) $\bar{\Pi}$ is rapidly varying at 0 (∞) with index $-\infty$, in which case $Y_k = 1$ a.s.

Theorem 2 has some important applications to the asymptotic distribution of the ratio of two consecutive ordered jumps $m_t^{(k+1)}/m_t^{(k)}$, $k \geq 1$, of a Lévy process. Let X_t , $t \geq 0$, be a Lévy process whose Lévy measure Λ is concentrated on $(0, \infty)$. Here in addition to $\bar{\Lambda}(x) \rightarrow 0$ as $x \rightarrow \infty$, we require that

$$\int_0^\infty \min\{1, x^2\} \Lambda(dx) < \infty. \quad (5)$$

In this setup one has the distributional representation for $k \geq 1$

$$(m_t^{(k)}, m_t^{(k+1)}) \stackrel{\mathcal{D}}{=} (\varphi(S_k/t), \varphi(S_{k+1}/t)), \quad (6)$$

with φ defined for $s > 0$ to be

$$\varphi(s) = \sup\{y : \bar{\Lambda}(y) > s\}. \quad (7)$$

It is readily checked that φ is nonincreasing and right continuous. Moreover, whenever Λ is the Lévy measure of a subordinator V_t , condition (1) holds, which is equivalent to

$$\int_\delta^\infty \varphi(s) ds < \infty, \text{ for any } \delta > 0. \quad (8)$$

The distributional representation in (6) follows from Proposition 1 in Kevei and Mason [7], see the proof of Theorem 1 below. For general spectrally positive Lévy processes it can be deduced using the same methods that Maller and Mason [9] derived the distributional representation for a Lévy process given in their Proposition 5.7.

When applying Theorem 2 to the asymptotic distribution of consecutive ordered jumps at 0 or ∞ of a Lévy processes X_t whose Lévy measure Λ is concentrated on $(0, \infty)$, we have to keep in mind that (5) must always hold and (1) must be satisfied whenever X_t is a subordinator. For instance in the case of a subordinator V_t , whenever $m_t^{(k+1)}/m_t^{(k)}$ converges in distribution to a random variable Y_k as $t \downarrow 0$, Theorem 2 says that $\bar{\Lambda}$ is regularly varying at 0. Further since (1) must hold, the parameter $-\alpha$ is necessarily in $[-1, 0]$, while there is no such restriction when considering convergence in distribution as $t \rightarrow \infty$. We note that in case of general Lévy processes for $k = 1$ the sufficiency part corresponding to part (ii) in Theorem 2 is given in Theorem 3.1 in [5].

In the special case when V_t is an α -stable subordinator, $\alpha \in (0, 1)$, and $m^{(1)} > m^{(2)} > \dots$ is its jump sequence on $[0, 1]$, then $(m^{(1)}/V_1, m^{(2)}/V_1, \dots)$ has the Poisson–Dirichlet law with parameter $(\alpha, 0)$ ($\text{PD}(\alpha, 0)$), see Bertoin [1] p. 90. The ratio of the $(k+1)^{\text{th}}$ and k^{th} element of a vector, which has the $\text{PD}(\alpha, 0)$ law, has the $\text{Beta}(k\alpha, 1)$ distribution (Proposition 2.6 in [1]).

2 Proofs

In the proofs we only consider the case when $t \downarrow 0$, as the $t \rightarrow \infty$ case is nearly identical.

2.1 Proof of Theorem 1

First we calculate the Laplace exponent of the ratio using the notation φ defined in (7). We see by the nonincreasing version of the change of variables formula stated in (4.9) Proposition of Revuz and Yor [10], which is given in Lemma 1 in [7],

$$\begin{aligned} \mathbf{E}e^{-\lambda V_t} &= \exp \left\{ -t \int_0^\infty \left(1 - e^{-\lambda v} \right) \Lambda(\mathrm{d}v) \right\} \\ &= \exp \left\{ -t \int_0^\infty \left(1 - e^{-\lambda \varphi(x)} \right) \mathrm{d}x \right\}. \end{aligned}$$

The key ingredient of our proofs is a distributional representation of the subordinator V_t given in Kevei and Mason (Proposition 1 in [7]), which follows from a general representation by Rosiński [11]. It states that for $t > 0$

$$V_t \stackrel{\mathcal{D}}{=} \sum_{i=1}^{\infty} \varphi \left(\frac{S_i}{t} \right). \quad (9)$$

From the proof of this result it is clear that $\varphi(S_i/t)$ corresponds to $m_t^{(i)}$, for $i \geq 1$. Therefore

$$\frac{V_t^{(k)}}{m_t^{(k+1)}} \stackrel{\mathcal{D}}{=} \frac{\sum_{i=k+1}^{\infty} \varphi(S_i/t)}{\varphi(S_{k+1}/t)}.$$

Conditioning on $S_{k+1} = s$ and using the independence we can write

$$\begin{aligned} \sum_{i=k+2}^{\infty} \varphi(S_i/t) &= \sum_{i=k+2}^{\infty} \varphi\left(\frac{s}{t} + \frac{S_i - s}{t}\right) \\ &\stackrel{\mathcal{D}}{=} \sum_{i=1}^{\infty} \varphi\left(\frac{s}{t} + \frac{S_i}{t}\right) \\ &= \sum_{i=1}^{\infty} \varphi_{s/t}(S_i/t), \end{aligned}$$

where $\varphi_y(x) = \varphi(y + x)$. Note that the latter sum has the same form as in (9), therefore it is equal in distribution to a subordinator $V^{(s/t)}(t)$ with Laplace transform

$$\begin{aligned} \mathbf{E} e^{-\lambda V_t^{(s/t)}} &= \exp \left\{ -t \int_0^{\infty} (1 - e^{-\lambda \varphi_{s/t}(x)}) dx \right\} \\ &= \exp \left\{ -t \int_{s/t}^{\infty} (1 - e^{-\lambda \varphi(x)}) dx \right\}. \end{aligned} \tag{10}$$

Now we can compute the Laplace transform of the ratio $V_t^{(k)}/m_t^{(k+1)}$. Since S_{k+1} has Gamma($k+1, 1$) distribution, the law of total probability and (10) give

$$\begin{aligned} \mathbf{E} \exp \left\{ -\lambda \frac{V_t^{(k)}}{m_t^{(k+1)}} \right\} &= \mathbf{E} \exp \left\{ -\lambda \frac{\sum_{i=k+1}^{\infty} \varphi(S_i/t)}{\varphi(S_{k+1}/t)} \right\} \\ &= \int_0^{\infty} \frac{s^k}{k!} e^{-s} \left[e^{-\lambda} \mathbf{E} \exp \left\{ -\frac{\lambda}{\varphi(s/t)} \sum_{i=1}^{\infty} \varphi_{s/t}(S_i/t) \right\} \right] ds \\ &= e^{-\lambda} \int_0^{\infty} \frac{s^k}{k!} e^{-s} \exp \left\{ -t \int_{s/t}^{\infty} [1 - e^{-\lambda \frac{\varphi(x)}{\varphi(s/t)}}] dx \right\} ds \\ &= \frac{t^{k+1}}{k!} e^{-\lambda} \int_0^{\infty} u^k \exp \left\{ -t \left(u + \int_u^{\infty} [1 - e^{-\lambda \frac{\varphi(x)}{\varphi(u)}}] dx \right) \right\} du \\ &= \frac{t^{k+1}}{k!} e^{-\lambda} \int_0^{\infty} u^k e^{-t\Psi(u, \lambda)} du, \end{aligned} \tag{11}$$

where

$$\Psi(u, \lambda) = u + \int_u^\infty \left[1 - e^{-\lambda \frac{\varphi(x)}{\varphi(u)}} \right] dx. \quad (12)$$

Since φ is right continuous on $(0, \infty)$, $\Psi(\cdot, \lambda)$ is also right continuous on $(0, \infty)$. Further a short calculation shows that this function is strictly increasing for any $\lambda > 0$, moreover for $u_1 > u_2$

$$\Psi(u_1, \lambda) - \Psi(u_2, \lambda) \geq e^{-\lambda}(u_1 - u_2).$$

Clearly $\Psi(\infty, \lambda) = \infty$ and therefore

$$\Psi_k(u, \lambda) := \Psi\left(((k+1)u)^{1/(k+1)}, \lambda\right)$$

has a right continuous increasing inverse function given by

$$Q_\lambda(s) = \inf \{v : \Psi_k(v, \lambda) > s\}, \text{ for } s \geq 0,$$

such that $Q_\lambda(0) = 0$ and $\lim_{x \rightarrow \infty} Q_\lambda(x) = \infty$. (For the right continuity part see (4.8) Lemma in Revuz and Yor [10].)

Necessity. Assuming that $V_t^{(k)}/m_t^{(k+1)}$ converges in distribution as $t \rightarrow 0$ to some extended random variable W_k , we can apply Theorem 2a on p. 210 of Feller [8] to conclude that its Laplace transform also converges, i.e.

$$\begin{aligned} \int_0^\infty u^k e^{-t\Psi(u, \lambda)} du &= \int_0^\infty e^{-t\Psi_k(v, \lambda)} dv \\ &= \int_0^\infty e^{-ty} dQ_\lambda(y) \sim \frac{e^\lambda g_k(\lambda) k!}{t^{k+1}}, \text{ as } t \rightarrow 0, \end{aligned}$$

where $g_k(\lambda) = \mathbf{E}e^{-\lambda W_k}$, and W_k can possibly have a defective distribution, i.e. possibly $\mathbf{P}\{W_k = \infty\} > 0$. (Here we used the change of variables formula given in (4.9) Proposition in Revuz and Yor [10].) By Karamata's Tauberian theorem (Theorem 1.7.1 in [2])

$$Q_\lambda(y) \sim \frac{y^{k+1}}{k+1} e^\lambda g_k(\lambda), \quad \text{as } y \rightarrow \infty,$$

and thus by Theorem 1.5.12 in [2]

$$\Psi_k(v, \lambda) \sim \left(\frac{(k+1)v}{e^\lambda g_k(\lambda)} \right)^{1/(k+1)}, \quad \text{as } v \rightarrow \infty,$$

and hence

$$\Psi(u, \lambda) \sim u \left[e^\lambda g_k(\lambda) \right]^{-\frac{1}{k+1}}, \quad \text{as } u \rightarrow \infty.$$

Substituting back into (12) we obtain for any $\lambda > 0$

$$\lim_{u \rightarrow \infty} \frac{1}{u} \int_u^\infty \left(1 - e^{-\lambda \frac{\varphi(x)}{\varphi(u)}} \right) dx = \left[e^\lambda g_k(\lambda) \right]^{-\frac{1}{k+1}} - 1. \quad (13)$$

Note that the limit W_k is ≥ 1 , with probability 1, and so $g_k(\lambda) \leq e^{-\lambda}$. Thus for any λ

$$\left[e^\lambda g_k(\lambda) \right]^{-\frac{1}{k+1}} - 1 \geq 0.$$

For any $x \geq 0$ we have $1 - e^{-x} \leq x$. Therefore by (13) we obtain for any $\lambda > 0$

$$\liminf_{u \rightarrow \infty} \frac{1}{u\varphi(u)} \int_u^\infty \varphi(x) dx \geq \frac{1}{\lambda} \left(\left[e^\lambda g_k(\lambda) \right]^{-\frac{1}{k+1}} - 1 \right). \quad (14)$$

On the other hand, by monotonicity $\varphi(x)/\varphi(u) \leq 1$ for $u \leq x$. Therefore for any $0 < \varepsilon < 1$ there exists a $\lambda_\varepsilon > 0$, such that for all $0 < \lambda < \lambda_\varepsilon$

$$1 - e^{-\lambda \frac{\varphi(x)}{\varphi(u)}} \geq (1 - \varepsilon) \frac{\lambda \varphi(x)}{\varphi(u)}, \quad \text{for } x \geq u.$$

Using again (13) and keeping (8) in mind, this implies that for such λ

$$\limsup_{u \rightarrow \infty} \frac{1}{u\varphi(u)} \int_u^\infty \varphi(x) dx \leq \frac{1}{1 - \varepsilon} \frac{1}{\lambda} \left(\left[e^\lambda g_k(\lambda) \right]^{-\frac{1}{k+1}} - 1 \right). \quad (15)$$

In particular, we obtain that, whenever $g_k(\lambda) \neq 0$ (i.e. $\mathbf{P}\{W_k < \infty\} > 0$)

$$0 \leq \liminf_{u \rightarrow \infty} \frac{1}{u\varphi(u)} \int_u^\infty \varphi(x) dx \leq \limsup_{u \rightarrow \infty} \frac{1}{u\varphi(u)} \int_u^\infty \varphi(x) dx < \infty.$$

Note that in (14) the greatest lower bound is 0 for all $\lambda > 0$ if and only if $g_k(\lambda) = e^{-\lambda}$, in which case $W_k = 1$. Then the upper bound for the limsup in (15) is 0, thus

$$\lim_{u \rightarrow \infty} \frac{1}{u\varphi(u)} \int_u^\infty \varphi(x) dx = 0,$$

which by Proposition 2.6.10 in [2] applied to the function $f(x) = x\varphi(x)$ implies that $\varphi \in \text{RV}_\infty(-\infty)$, and so, by Theorem 2.4.7 in [2], $\bar{\Lambda}$ is slowly varying at 0. We have proved that $W_k = 1$ if and only if $\bar{\Lambda}$ is slowly varying at 0.

In the following we assume that $\mathbf{P}\{W_k > 1\} > 0$, therefore the liminf in (14) is strictly positive. Let

$$a = \liminf_{\lambda \downarrow 0} \frac{1}{\lambda} \left(\left[e^\lambda g_k(\lambda) \right]^{-\frac{1}{k+1}} - 1 \right) \leq \limsup_{\lambda \downarrow 0} \frac{1}{\lambda} \left(\left[e^\lambda g_k(\lambda) \right]^{-\frac{1}{k+1}} - 1 \right) = b.$$

By (15) and (14), $a > 0$ and $b < \infty$. Moreover

$$b \leq \liminf_{u \rightarrow \infty} \frac{1}{u\varphi(u)} \int_u^\infty \varphi(x) dx \leq \limsup_{u \rightarrow \infty} \frac{1}{u\varphi(u)} \int_u^\infty \varphi(x) dx \leq a,$$

which forces

$$a = b = \lim_{u \rightarrow \infty} \frac{1}{u\varphi(u)} \int_u^\infty \varphi(x) dx = \lim_{\lambda \downarrow 0} \frac{1}{\lambda} \left(\left[e^\lambda g_k(\lambda) \right]^{-\frac{1}{k+1}} - 1 \right).$$

By Karamata's theorem (Theorem 1.6.1 (ii) in [2]) we obtain that φ is regularly varying at infinity with parameter $-a^{-1} - 1 =: -\alpha^{-1}$, so Λ is regularly varying with parameter $-\alpha$ at zero with $\alpha \in (0, 1)$.

Let us consider the case when $W_k = \infty$ a.s., that is $V_t^{(k)}/m_t^{(k+1)} \xrightarrow{\mathbf{P}} \infty$. All the previous computations are valid, with $g_k(\lambda) = \mathbf{E}e^{-\lambda\infty} \equiv 0$. Thus, from (14) we have

$$\lim_{u \rightarrow \infty} \frac{1}{u\varphi(u)} \int_u^\infty \varphi(x) dx = \infty.$$

From this, through the change of variables formula we obtain (3).

Sufficiency and the limit. Consider first the special case when $\varphi(x) = x^{-\frac{1}{\alpha}}$, $\alpha \in (0, 1)$. Then a quick calculation gives

$$\frac{1}{u} \int_u^\infty \left(1 - e^{-\lambda \frac{\varphi(x)}{\varphi(u)}} \right) dx = \alpha \int_0^1 \left(1 - e^{-\lambda y} \right) y^{-\alpha-1} dy.$$

By formula (13) for the Laplace transform of the limit we obtain (2).

The sufficiency can be proved by standard arguments for regularly varying functions. Using Potter bounds (Theorem 1.5.6 in [2]) one can show that for $\alpha \in (0, 1)$

$$\lim_{u \rightarrow \infty} \frac{1}{u} \Psi(u, \lambda) = 1 + \alpha \int_0^1 \left(1 - e^{-\lambda y} \right) y^{-\alpha-1} dy,$$

from which, through formula (11), the convergence readily follows. As already mentioned, cases (ii) and (iii) are treated in [5].

2.2 Proof of Theorem 2

Using that $\psi(y) \leq x$ if and only if $\bar{\Pi}(x) \leq y$, for the distribution function of the ratio we have for $x \in (0, 1)$

$$\begin{aligned}
\mathbf{P}\{r_k(t) \leq x\} &= \mathbf{P}\left\{\frac{\psi(S_{k+1}/t)}{\psi(S_k/t)} \leq x\right\} \\
&= \int_0^\infty \frac{s^{k-1}}{(k-1)!} e^{-s} \mathbf{P}\left\{\psi\left(\frac{s+S_1}{t}\right) \leq x\psi\left(\frac{s}{t}\right)\right\} ds \\
&= \int_0^\infty \frac{s^{k-1}}{(k-1)!} e^{-s} e^{-[t\bar{\Pi}(x\psi(s/t))-s]} ds \\
&= \frac{t^k}{(k-1)!} \int_0^\infty u^{k-1} e^{-t\bar{\Pi}(x\psi(u))} du.
\end{aligned} \tag{16}$$

Necessity. Assume that the limit distribution function G_k exists. Write

$$\frac{t^k}{(k-1)!} \int_0^\infty u^{k-1} e^{-t\bar{\Pi}(x\psi(u))} du = \frac{t^k}{(k-1)!} \int_0^\infty e^{-t\Phi_k(v,x)} dv, \tag{17}$$

where $\Phi_k(v, x) = \bar{\Pi}(x\psi((kv)^{1/k}))$. Note that for each $x \in (0, 1)$ the function $\Phi_k(\cdot, x)$ is monotone nondecreasing, since $\bar{\Pi}$ and ψ are both monotone nonincreasing. Let

$$\mathcal{G}_k = \{x : x \text{ is a continuity point of } G_k \text{ in } (0, 1) \text{ such that } G_k(x) > 0\}.$$

First assume that $\mathbf{P}\{Y_k < 1\} > 0$. Clearly we can now proceed as in the proof of Theorem 1 to apply Karamata's Tauberian theorem (Theorem 1.7.1 in [2]) to give that for any $x \in \mathcal{G}_k$,

$$\lim_{u \rightarrow \infty} \frac{\bar{\Pi}(x\psi(u))}{u} = [G_k(x)]^{-\frac{1}{k}}. \tag{18}$$

In fact, there is a small difference here compared to the proof of Theorem 1. We have to be more cautious, as $\Phi_k(v, x)$ is not necessarily right-continuous as a function of $v > 0$. To use the machinery from the proof of Theorem 1 we need to consider the right-continuous version $\tilde{\Phi}_k(v, x) := \Phi_k(v+, x)$. Since, in (17) we integrate with respect to the Lebesgue measure and Φ_k and $\tilde{\Phi}_k$ are equal almost everywhere, substituting Φ_k with $\tilde{\Phi}_k$ leaves the integral unchanged. Therefore, proceeding as before we obtain that

$$\tilde{\Phi}_k(v, x) \sim \left(\frac{kv}{G_k(x)}\right)^{1/k}, \quad \text{as } v \rightarrow \infty,$$

and since the right-hand function is continuous, we also get that

$$\Phi_k(v, x) \sim \left(\frac{kv}{G_k(x)} \right)^{1/k}, \quad \text{as } v \rightarrow \infty,$$

form which now (18) does indeed follow.

We claim that (18) implies the regular variation of $\bar{\Pi}$. When $\bar{\Pi}$ is continuous and strictly decreasing we get by changing variables to $\psi(u) = t$, $u = \bar{\Pi}(t)$, that we have for any $x \in \mathcal{G}_k$

$$\lim_{t \downarrow 0} \frac{\bar{\Pi}(tx)}{\bar{\Pi}(t)} = [G_k(x)]^{-\frac{1}{k}},$$

which by an easy application of Proposition 1.10.5 in [2] implies that $\bar{\Pi}$ is regularly varying.

Note that the jumps of $\bar{\Pi}$ correspond to constant parts of ψ , and vice versa. Put $\mathcal{J} = \{z : \bar{\Pi}(z-) > \bar{\Pi}(z)\}$ for the jump points of $\bar{\Pi}$. For $z \in \mathcal{J}$ and $y \in [\bar{\Pi}(z), \bar{\Pi}(z-))$ we have $\psi(y) = z$. Substituting into (18) we have

$$\lim_{z \downarrow 0, z \in \mathcal{J}} \frac{\bar{\Pi}(xz)}{\bar{\Pi}(z)} = [G_k(x)]^{-\frac{1}{k}}, \quad \text{and} \quad \lim_{z \downarrow 0, z \in \mathcal{J}} \frac{\bar{\Pi}(xz)}{\bar{\Pi}(z-)} = [G_k(x)]^{-\frac{1}{k}}. \quad (19)$$

To see how the second limit holds in (19) note that for any $0 < \varepsilon < 1$ and $z \in \mathcal{J}$, we have $\psi(\varepsilon \bar{\Pi}(z) + (1 - \varepsilon) \bar{\Pi}(z-)) = z$ and thus

$$\lim_{z \downarrow 0, z \in \mathcal{J}} \frac{\bar{\Pi}(xz)}{\varepsilon \bar{\Pi}(z) + (1 - \varepsilon) \bar{\Pi}(z-)} = [G_k(x)]^{-\frac{1}{k}}.$$

Since $0 < \varepsilon < 1$ can be chosen arbitrarily close to 0 this implies the validity of the second limit in (19). Therefore by choosing any $x \in \mathcal{G}_k$ we get

$$\lim_{z \downarrow 0} \frac{\bar{\Pi}(z-)}{\bar{\Pi}(z)} = 1. \quad (20)$$

Let

$$\mathcal{A} = \{z > 0 : \bar{\Pi}(z - \varepsilon) > \bar{\Pi}(z) \text{ for all } z > \varepsilon > 0\}.$$

This set contains exactly those points z for which $\psi(\bar{\Pi}(z)) = z$. With this notation formula (18) can be written as

$$\lim_{z \downarrow 0, z \in \mathcal{A}} \frac{\bar{\Pi}(xz)}{\bar{\Pi}(z)} = [G_k(x)]^{-\frac{1}{k}}, \quad \text{for } x \in \mathcal{G}_k. \quad (21)$$

This together with (20) will allow us to apply Proposition 1.10.5 in [2] to conclude that $\bar{\Pi}$ is regularly varying. We shall need the following technical lemma.

Lemma 1. *Whenever (20) holds, there exists a strictly decreasing sequence $z_n \in \mathcal{A}$ such that $z_n \rightarrow 0$ and*

$$\lim_{n \rightarrow \infty} \frac{\bar{\Pi}(z_{n+1})}{\bar{\Pi}(z_n)} = 1. \quad (22)$$

Proof. Choose $z_1 \in \mathcal{A}$ such that $\bar{\Pi}(z_1) > 0$, and define for each $n \geq 1$

$$z_{n+1} = \sup \left\{ z > 0 : \bar{\Pi}(z) > \left(1 + \frac{1}{n}\right) \bar{\Pi}(z_n-) \right\}.$$

Notice that the sequence $\{z_n\}$ is well-defined, since $\bar{\Pi}(0+) = \infty$ and it is decreasing. Further we have

$$\bar{\Pi}(z_{n+1}-) \geq \left(1 + \frac{1}{n}\right) \bar{\Pi}(z_n-) \text{ and } \bar{\Pi}(z_{n+1}) \leq \left(1 + \frac{1}{n}\right) \bar{\Pi}(z_n-),$$

where the second inequality follows by right continuity of $\bar{\Pi}$. Also note that $z_{n+1} < z_n$, since otherwise if $z_{n+1} = z_n$, then

$$\bar{\Pi}(z_{n+1}-) = \bar{\Pi}(z_n-) \geq \left(1 + \frac{1}{n}\right) \bar{\Pi}(z_n-),$$

which is impossible. Observe that each z_{n+1} is in \mathcal{A} since by the definition of z_{n+1} for all $0 < \varepsilon < z_{n+1}$

$$\bar{\Pi}(z_{n+1} - \varepsilon) > \left(1 + \frac{1}{n}\right) \bar{\Pi}(z_n-) \geq \bar{\Pi}(z_{n+1}).$$

Clearly since $\{z_n\}$ is a decreasing and positive sequence, $\lim_{n \rightarrow \infty} z_n = z^*$ exists and is ≥ 0 . By construction

$$\bar{\Pi}(z_{n+1}-) \geq \left(1 + \frac{1}{n}\right) \bar{\Pi}(z_n-) \geq \prod_{k=1}^n \left(1 + \frac{1}{k}\right) \bar{\Pi}(z_1-).$$

The infinite product $\prod_{n=1}^{\infty} (1 + 1/n) = \infty$ forces $z^* = 0$. Also by construction we have

$$1 \leq \frac{\bar{\Pi}(z_{n+1})}{\bar{\Pi}(z_n-)} = \frac{\bar{\Pi}(z_{n+1})}{\bar{\Pi}(z_n)} \left(\frac{\bar{\Pi}(z_n)}{\bar{\Pi}(z_n-)} \right) \leq 1 + \frac{1}{n}.$$

By (20) we have

$$\lim_{n \rightarrow \infty} \frac{\bar{\Pi}(z_n)}{\bar{\Pi}(z_n-)} = 1.$$

Therefore we get (22). \square

According to Proposition 1.10.5 in [2] to establish that $\bar{\Pi}$ is regularly varying at zero it suffices to produce λ_1 and λ_2 in $(0, 1)$ such that for $i = 1, 2$

$$\frac{\bar{\Pi}(\lambda_i z_n)}{\bar{\Pi}(z_n)} \rightarrow d_i \in (0, \infty), \text{ as } n \rightarrow \infty,$$

where $(\log \lambda_1) / (\log \lambda_2)$ is finite and irrational. This can clearly be done using (21) and $\mathbf{P}\{Y_k < 1\} > 0$. Necessarily $\bar{\Pi}$ has index of regular variation parameter $-\alpha \in (-\infty, 0]$. For $\alpha \in (0, \infty)$ the limiting distribution function has the form (4). In the case $\alpha = 0$, $\bar{\Pi}$ is slowly varying at 0 and we get that $G_k(x) = 1$ for $x \in (0, 1)$, i.e. $Y_k = 0$ a.s.

Now consider the case when $\mathbf{P}\{Y_k = 1\} = 1$, i.e. $G_k(x) = 0$ for any $x \in (0, 1)$. We once more use Theorem 1.7.1 in [2], with $c = 0$ this time, and as an analog of (18) we obtain

$$\lim_{u \rightarrow \infty} \frac{\bar{\Pi}(x\psi(u))}{u} = \infty.$$

This readily implies that

$$\lim_{z \downarrow 0, z \in \mathcal{A}} \frac{\bar{\Pi}(xz)}{\bar{\Pi}(z)} = \infty.$$

Moreover, the analogs of formula (19) also hold, i.e.

$$\lim_{z \downarrow 0, z \in \mathcal{J}} \frac{\bar{\Pi}(xz)}{\bar{\Pi}(z)} = \infty, \quad \text{and} \quad \lim_{z \downarrow 0, z \in \mathcal{J}} \frac{\bar{\Pi}(xz)}{\bar{\Pi}(z-)} = \infty.$$

(Note, however, that this does not imply (20).) Let $z \notin \mathcal{A}$, and define $z' = \inf\{v : v \in \mathcal{A}, v > z\}$. Clearly, $z' \downarrow 0$ as $z \downarrow 0$. If $z' \in \mathcal{A}$ then necessarily it is a jump point, $z' \in \mathcal{J}$, and $\bar{\Pi}(z'-) = \bar{\Pi}(z)$. Then

$$\frac{\bar{\Pi}(xz)}{\bar{\Pi}(z)} = \frac{\bar{\Pi}(xz)}{\bar{\Pi}(z'-)} \geq \frac{\bar{\Pi}(xz')}{\bar{\Pi}(z'-)},$$

and the latter tends to ∞ as $z \downarrow 0$. On the other hand, when $z' \notin \mathcal{A}$ it is simple to see that $\bar{\Pi}(z') = \bar{\Pi}(z)$ and $\bar{\Pi}(z' + \varepsilon) < \bar{\Pi}(z')$ for any $\varepsilon > 0$. Moreover, we can find $z < z'' \in \mathcal{A}$, such that $\bar{\Pi}(z) \leq \bar{\Pi}(z'') + 1 \leq 2\bar{\Pi}(z'')$ (we tacitly assumed that z is small enough). Thus

$$\frac{\bar{\Pi}(xz)}{\bar{\Pi}(z)} \geq \frac{\bar{\Pi}(xz)}{\bar{\Pi}(z'') + 1} \geq \frac{\bar{\Pi}(xz'')}{2\bar{\Pi}(z'')},$$

and the lower bound goes to ∞ as $z \downarrow 0$. Summarizing, we have proved that

$$\lim_{z \downarrow 0} \frac{\overline{\Pi}(xz)}{\overline{\Pi}(z)} = \infty,$$

for any $x \in (0, 1)$, that is, $\overline{\Pi}$ is rapidly varying at 0 with index $-\infty$.

Sufficiency. Assume that $\overline{\Pi}$ is regularly varying at 0 with index $-\alpha \in (-\infty, 0)$. Then its asymptotic inverse function ψ is regularly varying at ∞ with index $-1/\alpha$, therefore simply

$$r_k(t) = \frac{\psi(S_{k+1}/t)}{\psi(S_k/t)} \rightarrow \left(\frac{S_k}{S_{k+1}} \right)^{1/\alpha} \quad \text{a.s., as } t \downarrow 0,$$

which has the distribution G_k in (4). Assume now that $\overline{\Pi}$ is slowly varying at 0. Then $\psi \in \text{RV}_\infty(-\infty)$, therefore

$$r_k(t) = \frac{\psi(S_{k+1}/t)}{\psi(S_k/t)} \rightarrow 0 \quad \text{a.s., as } t \downarrow 0.$$

Finally, if $\overline{\Pi} \in \text{RV}_0(-\infty)$ then ψ is slowly varying at infinity, so

$$r_k(t) = \frac{\psi(S_{k+1}/t)}{\psi(S_k/t)} \rightarrow 1 \quad \text{a.s., as } t \downarrow 0,$$

and the theorem is completely proved.

Acknowledgement. PK was supported by the Hungarian Scientific Research Fund OTKA PD106181 and by the European Union and co-funded by the European Social Fund under the project ‘Telemedicine-focused research activities on the field of Mathematics, Informatics and Medical sciences’ of project number TÁMOP-4.2.2.A-11/1/KONV-2012-0073. DM thanks the Bolyai Institute for their hospitality while this paper was being written. The authors thank the referee for the useful comments and suggestions.

References

- [1] J. Bertoin. *Random Fragmentation and Coagulation Processes*. Cambridge University Press (2006).
- [2] N.H. Bingham, C.M. Goldie, J.L. Teugels. *Regular Variation*. Encyclopedia of Mathematics and its Applications, 27, Cambridge University Press, Cambridge (1987).

- [3] N.H. Bingham, J.L. Teugels. Conditions implying domains of attraction. Proceedings of the Sixth Conference on Probability Theory (Braşov, 1979), pp. 23–34, Ed. Acad. R.S. România, Bucharest (1981).
- [4] L. Breiman. On some limit theorems similar to the arc-sin law. *Teor. Verojatnost. i Primenen.* **10**, 351–360 (1965).
- [5] B. Buchmann, Y. Fan, R. Maller. Distributional representations and dominance of a Lévy process over its maximal jump processes. Available on arXiv: <http://arxiv.org/abs/1409.4050>.
- [6] D.A. Darling. The influence of the maximum term in the addition of independent random variables. *Trans. Amer. Math. Soc.* **73**, 95–107 (1952).
- [7] P. Kevei, D.M. Mason. Randomly Weighted Self-normalized Lévy Processes. *Stochastic Processes and their Applications* **123** (2), 490–522 (2013).
- [8] W. Feller. *An Introduction to Probability Theory and Its Applications* Vol. II, John Wiley and Sons, Inc., New York–London–Sydney (1966).
- [9] R.A. Maller, D.M. Mason. Small-time compactness and convergence behavior of deterministically and self-normalised Lévy processes. *Trans. Amer. Math. Soc.* **362**, 2205–2248 (2010).
- [10] D. Revuz, M. Yor. *Continuous Martingales and Brownian Motion*. Springer Verlag, Berlin (1991).
- [11] J. Rosiński. Series representation of Lévy processes from the perspective of point processes, in: O.E. Barndorff-Nielsen, T. Mikosch and S.I. Resnick (Eds), *Lévy Processes — Theory and Applications*, Birkhauser, Boston, 401–415 (2001).
- [12] J.L. Teugels. Limit theorems on order statistics. *Ann. Probab.* **9**, no. 5, 868–880 (1981).