A More General Maximal Bernstein-type Inequality

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Abstract

We extend a general Bernstein-type maximal inequality of Kevei and Mason (2011) for sums of random variables.

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1 Introduction

Let X_1, X_2, \ldots be a sequence of random variables, and for any choice of $1 \le k \le l < \infty$ we denote the partial sum $S(k,l) = \sum_{i=k}^{l} X_i$, and define $M(k,l) = \max\{|S(k,k)|, \ldots, |S(k,l)|\}$. It turns out that under a variety of assumptions the partial sums S(k,l) will satisfy a generalized Bernstein-type inequality of the following form: for suitable constants A > 0, a > 0, $b \ge 0$ and $0 < \gamma < 2$ for all $m \ge 0$, $n \ge 1$ and $t \ge 0$,

$$\mathbf{P}\{|S(m+1,m+n)| > t\} \le A \exp\left\{-\frac{at^2}{n+bt^{\gamma}}\right\}.$$
(1.1)

Kevei and Mason [2] provide numerous examples of sequences of random variables X_1, X_2, \ldots , that satisfy a Bernstein-type inequality of the form (1.1). They show, somewhat unexpectedly, without any additional assumptions, a modified version of it also holds for M(1+m, n+m) for all $m \ge 0$ and $n \ge 1$. Here is their main result.

Theorem 1.1. Assume that for constants A > 0, a > 0, $b \ge 0$ and $\gamma \in (0,2)$, inequality (1.1) holds for all $m \ge 0$, $n \ge 1$ and $t \ge 0$. Then for every 0 < c < a there exists a C > 0 depending only on A, a, b and γ such that for all $n \ge 1$, $m \ge 0$ and $t \ge 0$,

$$\mathbf{P}\{M(m+1,m+n) > t\} \le C \exp\left\{-\frac{ct^2}{n+bt^{\gamma}}\right\}.$$
(1.2)

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There exists an interesting class of Bernstein-type inequalities that are not of the form (1.1). Here are two motivating examples.

Example 1. Assume that X_1, X_2, \ldots , is a stationary Markov chain satisfying the conditions of Theorem 6 of Adamczak [1] and let f be any bounded measurable function such that $Ef(X_1) = 0$. His theorem implies that for some constants D > 0, $d_1 > 0$ and $d_2 > 0$ for all $t \ge 0$ and $n \ge 1$,

$$P\{|S_n(f)| \ge t\} \le D^{-1} \exp\left(-\frac{Dt^2}{nd_1 + td_2\log n}\right),$$
(1.3)

where $S_n(f) = \sum_{i=1}^n f(X_i)$, and D/d_1 is related to the limiting variance in the central limit theorem.

Example 2. Assume that X_1, X_2, \ldots , is a strong mixing sequence with mixing coefficients $\alpha(n), n \ge 1$, satisfying for some $d > 0, \alpha(n) \le \exp(-2dn)$. Also assume that $EX_i = 0$ and for some $M > 0, |X_i| \le M$, for all $i \ge 1$. Theorem 2 of Merlevède, Peligrad and Rio [4] implies that for some constant D > 0 for all $t \ge 0$ and $n \ge 1$,

$$P\{|S_n| \ge t\} \le D \exp\left(-\frac{Dt^2}{nv^2 + M^2 + tM(\log n)^2}\right),$$
 (1.4)

where $S_n = \sum_{i=1}^n X_i$ and $v^2 = \sup_{i>0} \left(Var(X_i) + 2\sum_{j>i} |cov(X_i, X_j)| \right)$.

The purpose of this note to establish the following extended version of Theorem 1.1 that will show that a maximal version of inequalities (1.3) and (1.4) also holds.

Theorem 1.2. Assume that there exist constants A > 0 and a > 0 and a sequence of nondecreasing non-negative functions $\{g_n\}_{n\geq 1}$ on $(0,\infty)$, such that for all t > 0 and $n \geq 1$, $g_n(t) \leq g_{n+1}(t)$ and for all $0 < \rho < 1$

$$\lim_{n \to \infty} \inf \left\{ \frac{t^2}{g_n(t) \log t} : g_n(t) > \rho n \right\} = \infty, \tag{1.5}$$

where the infimum of the empty set is defined to be infinity, such that for all $m \ge 0$, $n \ge 1$ and $t \ge 0$,

$$P\{|S(m+1,m+n)| > t\} \le A \exp\left\{-\frac{at^2}{n+g_n(t)}\right\}.$$
(1.6)

Then for every 0 < c < a there exists a C > 0 depending only on A, a and $\{g_n\}_{n \ge 1}$ such that for all $n \ge 1$, $m \ge 0$ and $t \ge 0$,

$$P\{M(m+1, m+n) > t\} \le C \exp\left\{-\frac{ct^2}{n+g_n(t)}\right\}.$$
(1.7)

Note that condition (1.5) trivially holds when the functions g_n are bounded, since the corresponding sets are empty sets. However, in the interesting cases g_n 's are not bounded, and in this case the condition basically says that $g_n(t)$ increases slower than t^2 .

Essentially the same proof shows that the statement of Theorem 1.2 remains true if in the numerator of (1.6) and (1.7) the function t^2 is replaced by a regularly varying function at infinity f(t) with a positive index. In this case the t^2 in condition (1.5) must be replaced by f(t). Since we do not know any application of a result of this type, we only mention this generalization.

Proof. Choose any 0 < c < a. We prove our theorem by induction on n. Notice that by the assumption, for any integer $n_0 \ge 1$ we may choose $C > An_0$ to make the statement true for all $1 \le n \le n_0$. This remark will be important, because at some steps of the proof we assume that n is large enough. Also since the constants A and a in (1.6) are independent of m, we can without loss of generality assume m = 0.

Assume the statement holds up to some $n \ge 2$. (The constant C will be determined in the course of the proof.)

Case 1. Fix a t > 0 and assume that

$$g_{n+1}(t) \le \alpha \, n,\tag{1.8}$$

for some $0 < \alpha < 1$ be specified later. (In any case, we assume that $\alpha n \ge 1$.) Using an idea of [5], we may write for arbitrary $1 \le k < n$, 0 < q < 1 and p + q = 1 the inequality

$$\begin{split} P\{M(1,n+1)>t\} \leq & P\{M(1,k)>t\} + P\{|S(1,k+1)|>pt\} \\ & + P\{M(k+2,n+1)>qt\}. \end{split}$$

Let

$$u = \frac{n + g_{n+1}(qt) - q^2 g_{n+1}(t)}{1 + q^2}$$

Note that $u \leq n-1$ if $0 < \alpha < 1$ is chosen small enough depending on q, for n large enough. Notice that

$$\frac{t^2}{u+g_{n+1}(t)} = \frac{q^2 t^2}{n-u+g_{n+1}(qt)}.$$
(1.9)

 Set

$$k = \begin{bmatrix} u \end{bmatrix}. \tag{1.10}$$

Using the induction hypothesis and (1.6), keeping in mind that $1 \le k \le n-1$, we obtain

$$P\{M(1, n+1) > t\} \leq C \exp\left\{-\frac{ct^2}{k+g_k(t)}\right\} + A \exp\left\{-\frac{ap^2t^2}{k+1+g_{k+1}(pt)}\right\} + C \exp\left\{-\frac{cq^2t^2}{n-k+g_{n-k}(qt)}\right\} \leq C \exp\left\{-\frac{ct^2}{k+g_{n+1}(t)}\right\} + A \exp\left\{-\frac{ap^2t^2}{k+1+g_{n+1}(pt)}\right\} + C \exp\left\{-\frac{cq^2t^2}{n-k+g_{n+1}(qt)}\right\}.$$
(1.11)

Notice that we chose k to make the first and third terms in (1.11) almost equal, and since by (1.10)

$$\frac{t^2}{k+g_{n+1}(t)} \le \frac{q^2 t^2}{n-k+g_{n+1}(qt)}$$

the first term is greater than or equal to the third. First we handle the second term in formula (1.11), showing that whenever $g_{n+1}(t) \leq \alpha n$,

$$\exp\left\{-\frac{ap^2t^2}{k+1+g_{n+1}(pt)}\right\} \le \exp\left\{-\frac{ct^2}{n+1+g_{n+1}(t)}\right\}.$$

For this we need to verify that for $g_{n+1}(t) \leq \alpha n$,

$$\frac{ap^2}{k+1+g_{n+1}(pt)} > \frac{c}{n+1+g_{n+1}(t)},$$
(1.12)

which is equivalent to

$$ap^{2}(n + 1 + g_{n+1}(t)) > c(k + 1 + g_{n+1}(pt)).$$

Using that

$$k = \lceil u \rceil \le u + 1 = 1 + \frac{1}{1 + q^2} \left[n + g_{n+1}(qt) - q^2 g_{n+1}(t) \right],$$

it is enough to show

$$n\left(ap^{2} - \frac{c}{1+q^{2}}\right) + ap^{2} - 2c$$
$$+ \left[g_{n+1}(t)ap^{2} - g_{n+1}(pt)c - \frac{c}{1+q^{2}}\left(g_{n+1}(qt) - q^{2}g_{n+1}(t)\right)\right] > 0$$

Note that if the coefficient of n is positive, then we can choose α in (1.8) small enough to make the above inequality hold. So in order to guarantee (1.12) (at least for large n) we only have to choose the parameter p so that $ap^2 - c > 0$, which implies that

$$ap^2 - \frac{c}{1+q^2} > 0 \tag{1.13}$$

holds, and then select α small enough, keeping mind that we assume $\alpha n \geq 1$ and $k \leq n-1$. Next we treat the first and third terms in (1.11). Because of the remark above, it is enough to handle the first term. Let us examine the ratio of $C \exp\{-ct^2/(k+g_{n+1}(t))\}$ and $C \exp\{-ct^2/(n+1+g_{n+1}(t))\}$. Notice again that since $u+1 \geq k$, the monotonicity of $g_{n+1}(t)$ and $g_{n+1}(t) \leq \alpha n$ implies

$$n+1-k \ge n-u = n - \frac{n+g_{n+1}(qt) - q^2g_{n+1}(t)}{1+q^2}$$
$$\ge \frac{q^2n - (1-q^2)g_{n+1}(t)}{1+q^2}$$
$$\ge n\frac{q^2 - \alpha(1-q^2)}{1+q^2}$$
$$=: c_1n.$$

At this point we need that $0 < c_1 < 1$. Thus we choose α small enough so that

$$q^2 - \alpha(1 - q^2) > 0. \tag{1.14}$$

Also we get using $g_{n+1}(t) \leq \alpha n$ the bound

$$(n+1+g_{n+1}(t))(k+g_{n+1}(t)) \le 2n^2(1+\alpha)^2 =: c_2n^2,$$

which holds if n large enough. Therefore, we obtain for the ratio

$$\exp\left\{-ct^2\left(\frac{1}{k+g_{n+1}(t)} - \frac{1}{n+1+g_{n+1}(t)}\right)\right\} \le \exp\left\{-\frac{cc_1t^2}{c_2n}\right\} \le e^{-1},$$

whenever $cc_1t^2/(c_2n) \ge 1$, that is $t \ge \sqrt{c_2n/(cc_1)}$. Substituting back into (1.11), for $t \ge \sqrt{c_2n/(cc_1)}$ and $g_{n+1}(t) \le \alpha n$ we obtain

$$P\{M(1, n+1) > t\}$$

$$\leq \left(\frac{2}{e}C + A\right) \exp\{-ct^2/(n+1+g_{n+1}(t))\} \leq C \exp\{-ct^2/(n+1+g_{n+1}(t))\},$$

where the last inequality holds for C > Ae/(e-2).

Next assume that $t < \sqrt{c_2 n/(cc_1)}$. In this case choosing C large enough we can make the bound > 1, namely

$$C \exp\left\{-\frac{ct^2}{n+1+g_{n+1}(t)}\right\} \ge C \exp\left\{-\frac{cc_2n}{cc_1n}\right\} = Ce^{-c_2/c_1} \ge 1,$$

if $C > e^{c_2/c_1}$.

Case 2. Now we must handle the case $g_{n+1}(t) > \alpha n$. Here we apply the inequality

$$P\{M(1, n+1) > t\} \le P\{M(1, n) > t\} + P\{|S(1, n+1)| > t\}.$$

Using assumption (1.6) and the induction hypothesis, we have

$$\begin{split} P\{M(1, n+1) > t\} &\leq C \exp\left\{-\frac{ct^2}{n+g_n(t)}\right\} + A \exp\left\{-\frac{at^2}{n+1+g_{n+1}(t)}\right\} \\ &\leq C \exp\left\{-\frac{ct^2}{n+g_{n+1}(t)}\right\} + A \exp\left\{-\frac{at^2}{n+1+g_{n+1}(t)}\right\}. \end{split}$$

We will show that the right side $\leq C \exp\{-ct^2/(n+1+g_{n+1}(t))\}$. For this it is enough to prove

$$\exp\left\{-ct^{2}\left(\frac{1}{n+g_{n+1}(t)}-\frac{1}{n+1+g_{n+1}(t)}\right)\right\} + \frac{A}{C}\exp\left\{-\frac{t^{2}(a-c)}{n+1+g_{n+1}(t)}\right\} \le 1.$$
(1.15)

Using the bound following from $g_{n+1}(t) > \alpha n$ and recalling that $\alpha n \ge 1$ and $0 < \alpha < 1$, we get

$$\frac{t^2}{(n+g_{n+1}(t))(n+1+g_{n+1}(t))} \ge \frac{\alpha^2 t^2}{(1+\alpha)(1+2\alpha)g_{n+1}(t)^2} =: c_3 \frac{t^2}{g_{n+1}(t)^2},$$

and

$$\frac{t^2(a-c)}{n+1+g_{n+1}(t)} \ge \frac{t^2}{g_{n+1}(t)} \frac{\alpha(a-c)}{1+2\alpha} =: \frac{t^2}{g_{n+1}(t)}c_4.$$

Choose $\delta > 0$ so small such that $0 < x \le \delta$ implies $e^{-cc_3x^2} \le 1 - \frac{cc_3}{2}x^2$. For $t/g_{n+1}(t) \ge \delta$ the left-hand side of (1.15) is less then

$$e^{-cc_3\delta^2} + \frac{A}{C},$$

which is less than 1, for C large enough.

For $t/g_{n+1}(t) \leq \delta$ by the choice of δ the left-hand side of (1.15) is less then

$$1 - \frac{cc_3}{2} \frac{t^2}{g_{n+1}(t)^2} + \frac{A}{C} \exp\left\{-\frac{t^2}{g_{n+1}(t)}c_4\right\},\,$$

which is less than 1 if

$$\frac{cc_3}{2}\frac{t^2}{g_{n+1}(t)^2} > \frac{A}{C}\exp\left\{-\frac{t^2}{g_{n+1}(t)}c_4\right\}.$$

By (1.5), for any $0 < \eta < 1$ and all large enough n, $g_{n+1}(t) 1 \{g_{n+1}(t) > \alpha n\} \le \eta t^2$, so that for all large n, whenever $g_{n+1}(t) > \alpha n$, we have

$$\frac{t^2}{g_{n+1}(t)^2} \ge t^{-2},$$

and again by (1.5) for all large n, whenever $g_{n+1}(t) > \alpha n$, $t^2/g_{n+1}(t) \ge (3/c_4) \log t$. Therefore for all large n, whenever $g_{n+1}(t) \alpha n$,

$$\exp\left\{-\frac{t^2}{g_{n+1}(t)}c_4\right\} \le t^{-3},$$

which is smaller than $t^{-2}\frac{Ccc_3}{2A}$, for t large enough, i.e. for n large enough. The proof is complete.

By choosing $g_n(t) = bt^{\gamma}$ for all $n \ge 1$ we see that Theorem 1.2 gives Theorem 1.1 as a special case. Also note that Theorem 1.2 remains valid for sums of Banach space valued random variables with absolute value $|\cdot|$ replaced by norm $||\cdot||$. Theorem 1.2 permits us to derive the following maximal versions of inequalities (1.3) and (1.4).

Application 1. In Example 1 one readily checks that the assumptions of Theorem 1.2 are satisfied with $A = D^{-1}$ and $a = D/d_1$

$$g_n(t) = \left(\frac{td_2}{d_1}\right)\log n.$$

We get the maximal version of inequality (1.3) holding for any 0 < c < 1 and all $n \ge 1$ and t > 0

$$P\left\{\left|\max_{1\le m\le n} S_n(f)\right| \ge t\right\} \le C \exp\left(-\frac{cDt^2}{nd_1 + td_2\log n}\right),\tag{1.16}$$

for some constant $C \ge D^{-1}$ depending on $c, D^{-1}, D/d_1$ and $\{g_n\}_{n \ge 1}$.

Application 2. In Example 2 one can verify that the assumptions of the Theorem 1.2 hold with A = D and $a = D/v^2$ and

$$g_n(t) = \frac{M^2}{v^2} + \left(\frac{tM}{v^2}\right) \left(\log n\right)^2,$$

which leads to the maximal version of inequality (1.4) valid for any 0 < c < 1 and all $n \ge 1$ and t > 0

$$P\left\{\max_{1\le m\le n}|S_m|\ge t\right\}\le C\exp\left(-\frac{cDt^2}{nv^2+M^2+tM\left(\log n\right)^2}\right)$$
(1.17)

for some constant $C \ge D$ depending on c, D/v^2 and $\{g_n\}_{n\ge 1}$. See Corollary 24 of Merlevède and Peligrad [3] for a closely related inequality that holds for all $n \ge 2$ and $t > K \log n$ for some K > 0.

Remark There is a small oversight in the published version of the Kevei and Mason paper. Here are the corrections that fix it.

1. Page 1057, line -9: Replace " $1 \le k \le n$ " by " $1 \le k < n$ ".

2. Page 1057, line -7: Replace this line with

 $\leq \mathbf{P} \{ M(1,k) > t \} + \mathbf{P} \{ S(1,k+1) > pt \} + \mathbf{P} \{ M(k+2,n+1) > qt \}.$

3. Page 1058: Replace " $k + bp^{\gamma}t^{\gamma}$ " by " $k + 1 + bp^{\gamma}t^{\gamma}$ " in equations (2.4) and (2.5), as well as in line -13.

4. Page 1058: Replace " $ap^2 - c$ " by " $ap^2 - 2c$ " in line -9.

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References

- R. Adamczak, A tail inequality for suprema of unbounded empirical processes with applications to Markov chains. Electron. J. Probab. 13 (2008), 1000–1034.
- [2] P. Kevei and D.M. Mason, A note on a maximal Bernstein inequality. Bernoulli 17 (2011), 1054–1062.
- [3] F. Merlevède and M. Peligrad, Rosenthal-type inequalities for the maximum of partial sums of stationary processes and examples. Ann. Probab. To appear.
- [4] F. Merlevède, M. Peligrad, M. and E. Rio, Bernstein inequality and moderate deviations under strong mixing conditions. In: High Dimensional Probability V: The Luminy Volume, C. Houdré, V. Koltchinskii, D. M. Mason and M. Peligrad, eds., (Beachwood, Ohio, USA: IMS, 2009), 273–292.
- [5] F.A. Móricz, R.J. Serfling and W.F. Stout, Moment and probability bounds with quasisuperadditive structure for the maximum partial sum. Ann. Probab. 10 (1982), 1032–1040.